Incident fractures in HIV-infected individuals: a systematic review and meta-analysis

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Objective(s): Some but not all studies indicate that individuals with HIV infection are at increased risk of fracture. We systematically reviewed the literature to investigate whether incidence of fracture (both overall and fragility) differs between individuals with and without HIV.

Design: Systematic review and meta-analysis.

Methods: Medline, Scopus, and the Cochrane Library databases for all studies ever published up to September 28, 2012 and electronically available conference abstracts from CROI, ASBMR, IAS, and AIDS were searched. All studies reporting incidence of all fracture and fragility fracture in HIV-infected adults were included. A random effects model was used to calculate pooled estimates of incidence rate ratios (IRR) for studies that presented data for HIV-infected and controls. For all studies, incidence rates of fracture and predictors of fracture among HIV-infected individuals were summarized.

Results: Thirteen eligible studies were analyzed, of which seven included controls. Nine studies reported all incident fractures and ten presented incident fragility fractures. The pooled IRR was 1.58 (95% CI: 1.25, 2.00) for all fracture and 1.35 (95% CI: 1.10, 1.65) for fragility fracture. Smoking, white race, and older age were consistent predictors for fragility fractures.

Conclusions: Our results indicate that HIV-infection is associated with a modest increase in incident fracture. Future research should focus on clarifying risk factors, designing appropriate interventions, and the long-term implications of this increased risk for an aging HIV-infected population.

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Keywords: bone, fracture, fracture incidence, fragility fracture, HIV

Introduction

Low bone mineral density (BMD) and decreased bone mass have been reported in HIV-infected men and women as well as children and adolescents [1–4]. Multiple factors appear to be involved, including effects of HIV-1 viral proteins, inflammatory cytokines and antiretroviral therapy (ART) on bone cells and bone turnover [3,5–8]. However, whether HIV infection and/or ART increase fracture incidence has not been clearly established. Two large database studies suggest that the prevalence of ICD-9 coded or self-reported fracture is higher among HIV-infected individuals as compared to the general population [9,10], particularly among older individuals [10]. In several large cohort studies, incidence of fracture among HIV infected individuals was compared to incidence in prospectively enrolled HIV-uninfected individuals [11,12], individuals within the same clinic...
system [13], or the general population [14,15]. Other cohorts have reported incidence and predictors of fracture limited to HIV-infected individuals [16–18]. Complicating the comparison of these studies and the interpretation of their results are their contrasting study samples and differing research methodologies. Nevertheless, as the HIV-infected population continues to age, a sophisticated understanding of the relative impact of HIV and/or ART on fracture risk is needed to better inform clinical practice and future research.

We conducted a systematic review to investigate if incidence of fracture (both overall and fragility) differs in individuals with and without HIV. The primary objective of the study was to conduct a meta-analysis to estimate an incidence rate ratio (IRR) of incident fractures and incident fragility fractures in HIV-infected compared to controls. A secondary goal was to compare and contrast the reported incidence rates of fracture among HIV-infected individuals and to summarize predictors of fracture from available studies.

**Methods**

**Search process**
We conducted a literature search of Medline, Scopus, and the Cochrane Library for all studies ever published of fracture incidence in HIV-infected individuals through September 28, 2012. Studies that were electronically published ahead of print publication during this time period were eligible for inclusion. Additionally, published abstracts from the Conference on Retroviruses and Opportunistic Infections (CROI; 1997–2012), the scientific meeting of the American Society for Bone and Mineral Research (ASBMR; 2000–2011), the Conference on HIV Pathogenesis and Treatment of the International AIDS Society (IAS; 2001–2011), and the International AIDS Conference (AIDS; 2002–2012) were searched online. Specific search terms are provided in Table 1.

**Outcome and data analysis**
The primary outcomes in this review were the all-fracture and fragility fracture incidence rate ratios (IRR) of the HIV-infected group compared with the control group. Secondary outcomes were the all-fracture and fragility fracture incidence rates (IR) in HIV-infected individuals. Fracture incidence rate was calculated by dividing the number of fractures by the period of risk for all included patients during the study period expressed as per 1000 person-years of follow-up. Exact confidence intervals were calculated for studies that did not provide a confidence interval but did provide the number of subjects (N) and person-years of follow-up (PY). Requests for additional data were made to all authors of eligible studies that did not present results for both all fractures and fragility fractures. A pooled estimate of the IRR comparing HIV-infected to controls for both overall fractures and fragility fractures was calculated. To account for heterogeneity between studies, a random effects

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**Table 1. Search terms used for systematic review.**

<table>
<thead>
<tr>
<th>Database or Conference</th>
<th>Years</th>
<th>Search Terms</th>
</tr>
</thead>
<tbody>
<tr>
<td>Scopus</td>
<td>Up to September 28, 2012</td>
<td>“HIV”</td>
</tr>
<tr>
<td>Cochrane Library</td>
<td>Up to September 28, 2012</td>
<td>“HIV”</td>
</tr>
<tr>
<td>Conference on Retroviruses and Opportunistic Infections (CROI)</td>
<td>1997–2012</td>
<td>“HIV”</td>
</tr>
</tbody>
</table>
model was used to calculate the combined effect sizes. Statistical heterogeneity was assessed using the $I^2$ coefficient. Funnel plots were not assessed for publication bias as there were fewer than ten studies [19]. Analyses were conducted using SAS version 9.1.3 (SAS Institute, Cary, North Carolina) and RevMan (Version 5.2, The Cochrane Collaboration, Oxford, UK).

Results

Selection of studies

Figure 1 shows results from the literature search and study selection process. We identified 179 articles from the Medline database, 338 from Scopus, and two from the Cochrane library. We also identified a total of 128 abstracts, including 39 from CROI, 44 from ASBMR, 29 from AIDS, and 16 from IAS. After exclusion of duplicate records and studies that did not meet our inclusion criteria, 36 articles remained and we further evaluated the full texts of these publications. Twenty-three studies were excluded because they did not contain fracture incidence data or because they did not report fracture incidence rates in person-years of observation. Authors of four studies were contacted for additional data, and all responded to data or clarification requests (i.e. providing either confidence intervals or numbers for all fractures and/or fragility fractures and person-years) [12–14,17].

We used the provided estimated person-years for the two groups in Grund et al. (7500 person-years) [20]. A total of 13 studies met all inclusion criteria. A manual search of references cited by these studies did not yield any additional eligible articles.

The study by Young et al. reported subgroup data for each year in the study, and different sampling weights were used for each year to produce population estimates and standard errors in the National Hospital Ambulatory Medical Care Survey of outpatient departments (NHAMCS-OPD). Therefore, we only used data from the most recent year (2006) containing information for both the HIV-infected group and the control group [14]. Both the CROI abstract by Volk et al. [21] and the manuscript by Le Re et al. [22] were included in the analyses. Since the abstract contained data for fracture at the spine and hip and the published manuscript contained data only for fractures at the hip, we utilized the spine data from the abstract and the hip data from the manuscript.

Study characteristics

Table 2 summarizes the baseline characteristics of the 13 studies, published between 2007 and 2012, included in this systematic review. The meta-analysis includes only studies, published between 2007 and 2012, included in this systematic review. The meta-analysis includes only studies identified incident fractures using the International Statistical Classification of Diseases (ICD) codes [13,15–17,20–22], while others relied on patient reports [11,12,18,23–25]. Young et al. used discharge summaries or patient self-report to identify fractures for HIV-infected individuals, and ICD codes recorded in charts or discharge records for controls [14]. Of the 10 studies that reported fragility fractures, four defined fragility fractures in terms of trauma (fall from standing height or less, low-energy trauma, atraumatic or low trauma, and inadequate trauma) [11,15,17,25], and six studies defined them by location (hip and vertebral, and either forearm or humerus) [12–14,16,21,22,24].

Among the seven studies with a control group, three studies utilized prospectively enrolled HIV-uninfected individuals [11,12,23]. Two studies selected controls from the same health care system as HIV-infected individuals: Womack et al. used controls matched by age, race, gender, and site from the Veterans Aging Cohort Study Virtual Cohort (VACS-VC) and Lo Re et al. used hepatitis C virus (HCV)-uninfected/HIV-uninfected Medicaid recipients as controls [13,21,22]. Two studies had general population controls: Hansen et al. linked data from the Danish HIV Cohort study with the national Danish Civil Registration System and Danish National Hospital Registry to obtain matched population controls, and Young et al. used data from NHAMCS-OPDs to obtain bone fracture rates for the general United States population [14,15]. Of the seven studies that included a control group, four studies reported all incident fractures [11,14,15,23] and six reported incident fragility fractures [11–15,21,22]. However, since we were unable to obtain person-years of observation for fragility fractures from Young et al., data from that study was not utilized for the fragility fracture analyses [14]. Thus, four studies were used to perform meta-analysis for all fractures and five for fragility fractures (Fig. 2a and b).

Four of the thirteen studies included only or mainly men [12,13,16,23] and one involved females only [11]. Not every study reported sex-specific fracture incidence data. Young et al. reported that HIV-infected women aged 25–54 years experienced more fractures at vertebra and femoral neck sites, while HIV-positive men aged 25–54 had more fractures at the wrist and vertebra, compared to controls [14]. The age means or medians of the study populations were between 35 and 55 years [11,12,14,16–18,25]. Nine studies had a majority identified as “white” or “Caucasian” [12,14–17,20,24,25]. In general, the mean and median BMIs of the study subjects were similar across the studies [11,18]. The percentage of
Literature search by **September 28, 2012** (n=647)
- Databases (n=519)
  - Medline (n=179)
  - Scopus (n=338)
  - Cochrane (n=2)
- Conference abstracts (n=128)*
  - CROI (n=39)
  - ASBMR (n=44)
  - IAS (n=16)
  - AIDS (n=29)

Duplicate records removed (n=214)

Potentially relevant studies identified for title and abstract review (n=433)

Abstracts excluded (n=397)
- Irrelevant (n=221)
- Not in English (n=34)
- Participants not relevant (n=41)
- No fracture data (n=101)

Full-text articles retrieved and reviewed (n=36)

Articles excluded (n=23)
- No fracture incidence data (n=15)
- No person-year information (n=8)

Studies included (n=13)

*Conferences included: Conference on Retroviruses and Opportunistic Infections (CROI, 1997-2012), the scientific meeting of the American Society for Bone and Mineral Research (ASBMR, 2000-2011), the Conference on HIV Pathogenesis and Treatment of the International AIDS Society (IAS; 2001-2011), and the International AIDS Conference (AIDS; 2002-2012)

Fig. 1. Literature search and study selection.
### Table 2. Characteristics of studies meeting inclusion criteria (n = 13).

<table>
<thead>
<tr>
<th>Author (Year)</th>
<th>Country (Cohort)</th>
<th>Control Group (N)</th>
<th>N (Total)</th>
<th>N (HIV+)</th>
<th>N (Controls)</th>
<th>Sex (% M)</th>
<th>HIV+ Race/Ethnicity (%)</th>
<th>HIV+ Age (years)</th>
<th>HIV+ BMI (kg/m²)</th>
<th>HIV+ ART status (%)</th>
<th>Determination of Fracture</th>
<th>Fracture Type and (Definition)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Arnsten et al. (2007)</td>
<td>USA</td>
<td>Y</td>
<td>526</td>
<td>317</td>
<td>209</td>
<td>100</td>
<td>63 B; 12 W; 23 H</td>
<td>54.7, 5 (mean, SD)</td>
<td>15% with BMI ≥ 30</td>
<td>87 (experienced)</td>
<td>Self-report</td>
<td>All Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Bedimo et al. (2012)</td>
<td>USA</td>
<td>Y</td>
<td>56,660</td>
<td>56,660</td>
<td>98</td>
<td>100</td>
<td>With Of: 57 W; Without: 45 W</td>
<td>40 (median)</td>
<td>44 (median)</td>
<td>69 (experienced)</td>
<td>ICD codes</td>
<td>All Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Collin et al. (2009)</td>
<td>France</td>
<td>N</td>
<td>1283</td>
<td>1,283</td>
<td>77</td>
<td>100</td>
<td>83.1°</td>
<td>36.2 (median)</td>
<td>22 (median)</td>
<td>100 (experienced)</td>
<td>Self-report</td>
<td>All Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Grund et al. (2009)</td>
<td>USA, Australia,</td>
<td>N</td>
<td>2,753 (VS), 2,752 (DC)</td>
<td>2,753 (VS), 2,720 (DC)</td>
<td>73</td>
<td>29.1 B; 21.1 H; 43.6 W; 6.2 O</td>
<td>43 (median)</td>
<td>25.9 (median)</td>
<td>95 (experienced)</td>
<td>ICD codes</td>
<td>All Fracture Type and (Definition)</td>
<td></td>
</tr>
<tr>
<td>Hansen et al. (2012)</td>
<td>Denmark</td>
<td>Y</td>
<td>31,836</td>
<td>3,506</td>
<td>26,530</td>
<td>76</td>
<td>80 W</td>
<td>–</td>
<td>–</td>
<td>78 (during study period)</td>
<td>ICD codes</td>
<td>All Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Hasse et al. (2011)</td>
<td>Switzerland</td>
<td>N</td>
<td>8,444</td>
<td>8,444</td>
<td>71</td>
<td>100</td>
<td>66.1 W; 8.4 B; 1.6 H; 2.3 A; 21.5 U</td>
<td>45, 39-51 (median, IQR)</td>
<td>23.5 (median)</td>
<td>94 (experienced)</td>
<td>Self-report</td>
<td>All Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Volk et al. (2011)</td>
<td>USA</td>
<td>Y</td>
<td>462,656</td>
<td>95,827</td>
<td>63</td>
<td>100</td>
<td>27 W; 44 B; 9 H; 20 O</td>
<td>39, 33-46 (median, IQR)</td>
<td>–</td>
<td>100 (experienced)</td>
<td>ICD codes</td>
<td>Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Walker Harris et al.</td>
<td>USA</td>
<td>Y</td>
<td>5,106</td>
<td>–</td>
<td>–</td>
<td>100</td>
<td>‘70%W</td>
<td>–</td>
<td>–</td>
<td>–</td>
<td>Self-report</td>
<td>Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Womack et al. (2011)</td>
<td>USA</td>
<td>Y</td>
<td>119,318</td>
<td>40,115</td>
<td>79,203</td>
<td>100</td>
<td>55% B/H</td>
<td>–</td>
<td>–</td>
<td>75 (at any time during follow-up)</td>
<td>ICD codes</td>
<td>Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Yin et al. (2010)</td>
<td>USA</td>
<td>Y</td>
<td>2,191</td>
<td>1,728</td>
<td>663</td>
<td>0</td>
<td>13.3 W; 56.3 B; 27.2 H; 3.2 O</td>
<td>40.4, 8.8 (mean, SD)</td>
<td>28.5 (mean)</td>
<td>65.6 (at index visit)</td>
<td>Self-report</td>
<td>All Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Yin et al. (2012)</td>
<td>USA</td>
<td>N</td>
<td>4,640</td>
<td>4,640</td>
<td>83</td>
<td>100</td>
<td>48 W; 28.7 B; 20.4 H; 1.8 A; 1.2 O</td>
<td>39, 33-45 (median, IQR)</td>
<td>25, 22-28 (median, IQR)</td>
<td>26.5 (experienced at index)</td>
<td>Self-report</td>
<td>All Fracture Type and (Definition)</td>
</tr>
<tr>
<td>Yong et al. (2011)</td>
<td>Australia</td>
<td>N</td>
<td>2,424</td>
<td>2,424</td>
<td>89</td>
<td>100</td>
<td>Cases: 92 W; 3 B; 5 A; Cases: 49.5; Controls: 49.5 (mean)</td>
<td>Cases: 22.5 Controls: 25.2</td>
<td>Cases: 90 Controls: 80 (at time of fracture)</td>
<td>ICD codes</td>
<td>All Fracture Type and (Definition)</td>
<td></td>
</tr>
<tr>
<td>Young et al. (2011)</td>
<td>USA</td>
<td>Y</td>
<td>224,488,004</td>
<td>3,004</td>
<td>224,485,000</td>
<td>78</td>
<td>52 W; 33 B; 11.7 H; 3.5 O</td>
<td>40, 36-46 (median, IQR)</td>
<td>24.4 (median)</td>
<td>73 (experienced)</td>
<td>Discharge summary or self-report for HIV+ ICD codes for HIV-</td>
<td>All Fracture Type and (Definition)</td>
</tr>
</tbody>
</table>

A, Asian; B, Black; BMI, Body mass index; H, Hispanic/Latino; NA, not available; O, Other; U, unknown; W, White.

*Unspecified. 83.1% born in France or England.

*Not reported in paper [33].

*Not reported in paper [34].

*2006 data from Young presented, baseline characteristics are from all participants followed during 2000-2008 [14].
HIV-infected individuals on ART varied widely across the studies, from 27 to 100. While most studies classified those with any ART exposure as ART treated, Bedimo et al. applied a minimum exposure time of one month [16].

Fracture incidence

Figure 2a presents estimates of all fracture IRR in HIV-infected individuals compared to controls. The pooled estimate of the crude IRR for all fracture was 1.58 (95% CI: 1.25, 2.00). Figure 2b presents estimates of fragility fracture IRR in HIV-infected individuals compared to controls. The pooled estimate of the crude IRR for fragility fracture was 1.35 (95% CI: 1.10, 1.65). The assessment for heterogeneity was not significant for all fractures ($Q = 6.97, P = 0.07, I^2 = 57\%$) but was significant for fragility fractures ($Q = 44.58, P < 0.00001, I^2 = 91\%$). No significant changes in heterogeneity with respect to fragility fracture were noted when the analyses were repeated omitting either Yin et al. which included only women and prospectively enrolled HIV-uninfected controls or Walker-Harris et al. which included prospectively enrolled HIV-uninfected controls or both studies (results not shown) [11,12]. Table 3 presents estimates of all fracture and fragility fracture IR (per 1000 person-years) for HIV-infected individuals. Of note, there was greater variation in the crude incidence rates of all fractures, which ranged from 0.8 to 30.6 per 1000 person-years compared to the crude incidence rates of fragility fractures, which ranged from 1.4 to 7.4 per 1000 person-years.

Predictors of fracture

A summary of significant predictors of all fractures and fragility fractures in multivariable or adjusted analyses for HIV-infected individuals is presented in Table 3. Though predictors varied, several traditional risk factors remained consistent across studies, including older age [11,13–16], white race [11,13,15,16,23], low weight or BMI [13,14,16], smoking [11,13,15,16,24], and alcohol or substance abuse [13,14,18], particularly for fragility fractures. Diabetes, liver disease and a co-morbidity index were also been reported as significant predictors of fracture [11,13–15]. Although not a traditional risk factor, HCV was found to be a significant predictor of fracture in many studies [11,14,15,18,21,22]. Use of glucocorticoids was also found to be a significant predictor of fracture in three studies [13,17,24], while only one study found proton pump inhibitors to be associated with fractures [13]. As for HIV-specific variables, two studies found an association between low CD4 count and fracture risk [14,17]. Few studies reported an independent effect of ART types or exposure on fractures. Bedimo et al. found that tenofovir and ritonavir-boosted protease inhibitors (PIs) were associated with an increased risk of osteoporotic fractures,
Table 3. ???.

2: Crude incidence rates of all fractures and significant predictors of all fractures (in adjusted analyses) in HIV-infected individuals across 9 studies

<table>
<thead>
<tr>
<th>Author (Year)</th>
<th>All Fracture Crude Incidence Rate (95% CI) per 1000 person-years</th>
<th>Significant Predictors of All Fractures^a</th>
</tr>
</thead>
<tbody>
<tr>
<td>Arnsten et al. (2007)</td>
<td>30.6 (19.0, 46.8)</td>
<td>Non-black race, Low bone mineral density</td>
</tr>
<tr>
<td>Collin et al. (2009)</td>
<td>3.3 (2.0, 4.6)</td>
<td>Substance abuse, HCV co-infection</td>
</tr>
<tr>
<td>Grund et al. (2009)</td>
<td>0.8 (0.5, 1.41)</td>
<td>HCV co-infection^b</td>
</tr>
<tr>
<td>Hansen et al. (2012)</td>
<td>21.0 (19.8, 22.2)</td>
<td>Older age</td>
</tr>
<tr>
<td>Hasse et al. (2011)</td>
<td>7.1 (6.1, 8.3)</td>
<td>White race, Older age, Smoking, History of AIDS defining illness, Cumulative NNRTI use</td>
</tr>
<tr>
<td>Yin et al. (2010)</td>
<td>18.1 (15.4, 21.3)</td>
<td>Smoking, HCV co-infection, Bisphosphonate use, Steroid use</td>
</tr>
<tr>
<td>Yin et al. (2012)</td>
<td>4.0 (3.3, 4.8)</td>
<td>N/A</td>
</tr>
<tr>
<td>Yong et al. (2011)</td>
<td>5.3 (4.3, 6.5)</td>
<td>Substance abuse, HCV co-infection, Older age, CD4 count &lt;200 cells/ml, diabetes</td>
</tr>
<tr>
<td>Young et al. (2011)</td>
<td>8.7 (5.9, 12.7)</td>
<td></td>
</tr>
</tbody>
</table>

b: Crude incidence rates of fragility fractures and significant predictors of fragility fractures (in adjusted analyses) in HIV-infected individuals across 10 studies

<table>
<thead>
<tr>
<th>Author (Year)</th>
<th>Fragility Fracture Crude Incidence Rate (95% CI) per 1000 person-years</th>
<th>Significant Predictors of Fragility Fractures^a</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bedimo et al. (2012)</td>
<td>3.1 (2.9, 3.3)</td>
<td>White race, Older age, Tobacco use, BMI &lt;20, HCV co-infection, tenofovir/PI use in HAART era</td>
</tr>
<tr>
<td>Hansen et al. (2012)</td>
<td>7.4 (6.7, 8.2)</td>
<td>White race, Older age, Smoking, Medium co-morbidity Score</td>
</tr>
<tr>
<td>Hasse et al. (2011)</td>
<td>1.9 (1.7, 2.2)</td>
<td>Older age, Low CD4 count</td>
</tr>
<tr>
<td>Volk et al. (2011)</td>
<td>4.2 (4.0, 4.5) - HCV only</td>
<td>HCV co-infection^b</td>
</tr>
<tr>
<td>Walker Harris et al. (2012)</td>
<td>6.6 (6.3, 7.0) - HCV/HCV</td>
<td></td>
</tr>
<tr>
<td>Womack et al. (2011)</td>
<td>1.4 (1.0, 1.8)</td>
<td>Age^b</td>
</tr>
<tr>
<td>Yin et al. (2010)</td>
<td>2.5 (2.3, 2.7)</td>
<td>White race, Older age, Smoking, Low BMI, Alcohol abuse, Liver disease, Corticosteroid use, Proton pump inhibitor use, PI use</td>
</tr>
<tr>
<td>Yin et al. (2012)</td>
<td>5.8 (4.3, 7.7)</td>
<td>N/A</td>
</tr>
<tr>
<td>Young et al. (2011)</td>
<td>1.0 (0.7, 1.5)</td>
<td>N/A</td>
</tr>
<tr>
<td>Yong et al. (2011)</td>
<td>4.1 (3.2, 5.2)</td>
<td>Index CD4 count &lt;200/ml, steroid use, anti-epileptic medications</td>
</tr>
<tr>
<td>Young et al. (2011)</td>
<td>N/A</td>
<td>Older age, HCV co-infection, BMI&lt;18.5</td>
</tr>
</tbody>
</table>

BMI, body mass index; HCV, Hepatitis C Virus; NNRTI, non-nucleoside reverse transcriptase inhibitor; PI, protease inhibitor.

^aSignificant factors in multivariate models.

^bSignificant factor in stratified analysis.
adjusted for traditional risk factors among those who entered the cohort in the HAART era (1996–2009) [16].

Discussion

We found that HIV-infected individuals have a modestly increased risk for all fractures and fragility fractures compared to uninfected individuals or the general population. The test for statistical heterogeneity was not significant for the pooled estimate of all fractures; therefore, the IRR for all fractures appears to be a valid estimate. The test for heterogeneity, however, was significant for the pooled risk estimate for fragility fracture and therefore this IRR should be interpreted with caution. It should be noted that the overall IR of fracture among HIV-infected individuals varied widely across studies, possibly due to the differences in demographics of the study populations.

Of the four studies that reported IRRs of all fractures, those with mainly Caucasian subjects reported a risk increase while the two that did not reach significance had a majority of black participants [11,14,15,23]. With respect to fragility fracture, LoRe et al. and Womack et al. found an increased risk in primarily minority cohorts; the largest risk however was reported by Hansen et al. in which 80% of the subjects were Caucasian [13,15,21,22]. Non-black race is a well-known risk factor for osteoporosis. Our analysis suggests that race modifies the relationship between HIV infection and fracture and may explain some of the heterogeneity found in our analysis.

Review of the adjusted analyses of each study revealed that a number of traditional risk factors are importantly associated with fractures among individuals with HIV including smoking [11,13,15], use of glucocorticoids [17] or proton pump inhibitors [13], alcohol or substance abuse [17,18], low weight or BMI [13], and comorbidities such as diabetes and liver disease [11,13–15]. BMI may be of particular importance since a meta-analysis conducted by Boll and et al. found that lower body weight mediated the effect of HIV status on low BMD [26]; however, reporting of weight or BMI in available fracture studies was inconsistent and therefore, could not be evaluated in an adjusted analyses. Low body weight may not only increase risk of fracture by leading to decreased BMD, but may also be associated with increased prevalence of the frailty phenotype at earlier ages and predisposition to falls. While early frailty has been previously documented in HIV-infected men, current data on fall rates in HIV-infected individuals is limited [27,28] and is an important area for future research.

The role of HIV-specific factors in bone loss and fracture also remains unclear. Lower CD4 counts before ART initiation has been reported to be a predictor of both bone loss after ART initiation and increases in bone turnover [29]. Low current CD4 count (<200 cells/µl) in studies by Yong et al. [17] and Hasse et al. [25] were associated with increased fragility fracture incidence in multivariate analyses. Similarly, nadir CD4 count <200 cells/µl in Yong et al. was associated with increased all fracture incidence in multivariate analysis [14]. However, pre-treatment CD4 count was not associated with fracture incidence in the longitudinal follow-up of participants enrolled in randomized clinical trials of ART initiation [24], and nadir or current CD4 count was not predictive of fractures in other studies [13,15,25].

Several studies analyzed the association between ART exposure and fracture risk. Hansen et al. found that ART-exposed patients had a higher risk of fragility fracture than the general population that remained significant after adjusting for the comorbidity index (IRR = 1.6; 95%CI: 1.36–1.87), but ART-exposure was not associated with an increased risk of non-fragility fracture [15]. In a larger database of HIV-infected subjects, when analyses were restricted to subjects entering the cohort in the era of potent ART, Bedimo et al. found that exposure to either tenofovir or PIs was associated with increased incidence of osteoporotic fracture [16]. Most other studies failed to find an association between ART and fracture incidence; however, this may be partly due insufficient sample size and differences in the classification of antiretroviral regimens used across the studies.

Interestingly, HCV was consistently identified as an independent risk factor for both fragility and non-fragility incident fractures [11,14–16,18,21,22]. The increased risk of fracture is approximately 1.5–2 times greater in HIV/HCV co-infected than HIV mono-infected individuals [15,16,21,22]. One study found that BMD was lower in HIV/HCV co-infected than HIV mono-infected women, but not in men [30]. A recent study found that DXA derived hip geometry measures (buckling ratio and centroid position) differed between HIV/HCV co-infected men and uninfected controls, suggesting that fracture risk in HIV/HCV co-infected individuals may be due to compromised bone strength as well as lower bone density [31]. Although the mechanisms are not well understood, HIV/HCV co-infection appears to have a negative effect on bone strength and fracture risk. The relationship between HIV/HCV co-infection and fracture risk warrants further study.

Our study has a number of limitations. We cannot exclude the possibility of publication bias, though we attempted to minimize this by reviewing abstracts from major scientific meetings. We were also unable to evaluate funnel plots due to the small number of published studies that met our criteria for inclusion. As mentioned, there was considerable diversity across the studies in terms of
study populations and how fractures were determined thus introducing heterogeneity into our pooled analyses. The accuracy and reliability of the methods used to distinguish fragility from non-fragility fractures is also unknown, allowing for potential misclassification bias. In addition, we were not able to perform analyses on adjusted estimates of IR for important factors such as age, race, sex, weight, antiretroviral exposure and CD4 since this data was not available in most studies and those that conducted adjusted analyses included different variables.

This systematic review and meta-analysis suggests that HIV infection and/or ART are associated with modest increases in fracture risk with a pooled IRR of 1.58 (95% CI: 1.25, 2.00) for all fracture. Our results show the importance of including fracture outcomes in prospective studies that provide individual patient level data in order to establish the attributable risk of HIV infection and/or ART on fracture, as well as the contribution of other risk factors, including age, sex, race, weight and smoking. Validation of fracture prediction algorithms, such as FRAX [32], for HIV-infected individuals remains an important area of research. While at present the overall increase in risk of fracture is modest, this can be expected to increase in the future as the HIV-infected population continues to age, and studies of risk reduction interventions are warranted.

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Conflicts of interest

None declared.

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References


