

Association of Nucleos(t)ide Analogue Therapy With Reduced Risk of Hepatocellular Carcinoma in Patients With Chronic Hepatitis B—A Nationwide Cohort Study

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BACKGROUND & AIMS: Treatment for hepatitis B virus infection reduces the risk of hepatocellular carcinoma (HCC). However, the long-term protective effects for subgroups of patients with chronic hepatitis B are unclear. **METHODS:** We conducted a retrospective nationwide cohort study using data from Taiwan's National Health Insurance Research Database (from January 1, 1997, through December 31, 2010). Cumulative incidences were calculated and multivariable analyses were carried out after adjusting for competing mortality. Propensity scores were used to match 21,595 patients with chronic hepatitis B who received nucleoside analogue therapy for at least 90 days (treated cohort) with 21,595 untreated patients with chronic hepatitis B (controls), who received hepatoprotectants for at least 90 days. Data were collected from the treated cohort for a mean period of 3.46 years and from controls for 5.24 years. **RESULTS:** The treated cohort had a significantly lower 7-year incidence of HCC (7.32%; 95% confidence interval [CI], 6.77%–7.87%) than controls (22.7%; 95% CI, 22.1%–23.3%; $P < .001$). After adjusting for competing mortality and other confounders, nucleos(t)ide analogue treatment was associated with a reduced risk of HCC, with an adjusted hazard ratio of 0.37 (95% CI, 0.34–0.39; $P < .001$). Sensitivity analyses confirmed the association between nucleos(t)ide analogue treatment and reduced risk of HCC. Age, sex, cirrhosis, and diabetes mellitus modified this association. **CONCLUSIONS:** Based on a retrospective, nationwide study in Taiwan, nucleoside analogue therapy use is associated with reduced risk of HCC in patients with chronic hepatitis B virus infection.

Keywords: Hepatoma; HBV; Antiviral Agent; NHIRD.

Hepatocellular carcinoma (HCC) is the fifth most common cancer worldwide and the third most common cause of cancer-related mortality. It is estimated

that >748,000 new HCC cases and about 700,000 deaths occur annually worldwide.^{1,2} Many risk factors contribute to the development of HCC, including hepatitis B virus (HBV), hepatitis C virus, alcoholic liver disease, nonalcoholic steatohepatitis, and metabolic syndrome.³ Globally, and in Asia particularly, chronic HBV infection is the most frequent underlying cause of HCC and accounts for approximately half of HCC cases.^{4,5} Vaccines against HBV have successfully reduced the incidence of HBV in the younger generations, however, there are still >350 million patients infected with HBV worldwide.⁶ Chronic hepatitis B (CHB) infection not only causes hepatitis, but also leads to hepatic decompensation, cirrhosis, and hepatocellular carcinoma (HCC).^{1,3,6,7} HBV replication has been identified as a major element of immune-mediated liver tissue injury and disease progression. Higher HBV DNA levels are associated with increased risk of HCC development and recurrence.^{4,8,9}

Nucleos(t)ide analogue therapy effectively suppresses HBV replication by inhibiting HBV polymerase.^{10,11} Treatment with nucleos(t)ide analogues has been reported to delay disease progression in CHB patients.^{12,13} Regression of liver cirrhosis has been observed with long-term use of nucleos(t)ide analogues.^{14,15} In addition, nucleos(t)ide analogue therapy has been reported to be associated with reduced risk of HCC development and recurrence.^{16–18} In a meta-analysis that pooled 5 studies and a total of 2289 CHB patients, the risk of HCC was reduced by 78% among those receiving nucleos(t)ide analogue therapy relative to controls that did not receive nucleos(t)ide analogue.¹⁹ In another systematic review of 21 studies involving 3881

Abbreviations used in this paper: CHB, chronic hepatitis B; CI, confidence interval; HBV, hepatitis B virus; HCC, hepatocellular carcinoma; HR, hazard ratio; ICD-9, International Classification of Diseases, 9th Revision; NHI, National Health Insurance; NHIRD, National Health Insurance Research Database; NNT, number needed to treat.

treated and 534 untreated patients, the protective roles of nucleos(t)ide analogues in the development of HCC were confirmed.²⁰ In the CALM study (Cirrhosis Asian Lamivudine Multicentre Study), nucleos(t)ide analogue therapy attenuated 51% risk of HCC in patients with CHB.¹² However, some literature report no beneficial effects for nucleos(t)ide analogue in reducing HCC risk among CHB patients. In a recent cohort study, the risk of HCC development in patients with oral antiviral therapy is still significantly higher than in patients with inactive CHB.²¹ In another study, oral nucleos(t)ide analogue reduces the incidence of cirrhosis and risk of complications, but not development of HCC in cirrhotic patients.²² With this controversial evidence, a long-term population-based nationwide study will be helpful to investigate the association between nucleos(t)ide analogue treatment and risk of HCC in HBV-infected patients. In the present study, we conducted a nationwide cohort study to examine whether nucleos(t)ide analogue use is associated with reduced risk of HCC in CHB patients. In addition, we examined the number needed to treat (NNT) for 1 less HCC development.

Methods

Study Design

We conducted this nationwide cohort study based on Taiwan's National Health Insurance Research Database (NHIRD). The NHIRD was set up in 1997 when the National Health Insurance (NHI) program, a compulsory universal health insurance program for nearly all 23.7 million residents in Taiwan, was established. Comprehensive health care information, including diagnoses, prescriptions, and laboratory check-up items can be retrieved from the NHIRD, which has been described in detail in our previous studies.^{18,23,24} International Classification of Diseases, 9th Revision (ICD-9) codes are used to define diseases in this database. The accuracy of diagnosis of major diseases, such as ischemic stroke and acute coronary syndrome, has been validated.^{25,26} This study has been approved by the ethical review board of the National Health Research Institutes, Taiwan.

Study Subjects

CHB patients in the NHIRD were defined as those meeting the following 2 criteria: diagnosed with CHB (ICD-9 codes: 070.2, 070.3, and V02.61) 3 times in outpatient clinic or admitted with a diagnosis of HBV infection between January 1, 1997 and December 31, 2010 and having used nucleos(t)ide analogues or hepatoprotective agents (eg, silymarin, liver hydrolysate, and choline bitartrate). These CHB patients were followed up from the time of diagnosis of CHB until development of HCC, death, or December 31, 2010. We first excluded patients not using nucleos(t)ide analogues or hepatoprotective agents for at least 90 days. Nucleos(t)ide analogues have been covered under the NHI program for CHB patients since October 1, 2003. However, reimbursement for nucleos(t)ide analogues requires patients to fulfill certain criteria, such as twice-elevated serum alanine aminotransferase ([ALT] $\geq 2\times$) and elevated HBV DNA titer (>2000 IU/mL). The reimbursement duration ranges from 18 months to 36 months.¹⁸ The reimbursement for hepatoprotective agents requires only elevated ALT level (ie, ALT

$\geq 1\times$). Hepatoprotective agents have been reimbursed since the beginning of Taiwan's NHI program in 1997. Next, we excluded patients with hepatitis C virus (ICD-9 codes: 070.41, 070.44, 070.51, 070.54, and V02.62), human immunodeficiency virus (ICD-9 code: 042), other viral hepatitis (ICD-9 code: V02.69), and malignant tumors (ICD-9 codes: 104–208).

Study Cohorts

Among the eligible patients, the treated patients were defined as those who received nucleos(t)ide analogues for at least 90 days. Those who did not use nucleos(t)ide analogues for at least 90 days were defined as nontreated patients. Demographic data of these 2 groups were first compared. Then, propensity scores for estimating the probability of receiving nucleos(t)ide analogue therapy were developed using the logistic regression model to estimate the differences in baseline characteristics between the treated patients and the untreated patients. The propensity score approach used in our study was described in Dehejia et al.²⁷ We matched each treated patient with 1 untreated patient, based on propensity scores. Histograms before and after matching were examined to assess the success of the propensity scores in balancing the 2 groups. The index date of follow-up was the first date of nucleos(t)ide analogue prescription for the treated cohort and the first date of hepatoprotective agent prescription for the untreated cohort.

Hepatocellular Carcinoma Risk Analysis

HCC diagnosis was defined according to the major diagnosis of admission (ICD-9 codes: 155.0 and 155.2) or enrollment in the Registry for Catastrophic Illness Patient Database, a subset of the NHIRD. Patients were registered in the Registry for Catastrophic Illness Patient Database if their diagnoses were confirmed by pathology reports or typical imaging presentations. The first date of admission or enrollment in the Registry for Catastrophic Illness Patient Database was defined as the date of HCC development. We excluded patients with a diagnosis of HCC in the first 90 days after start of nucleos(t)ide analogue therapy or hepatoprotective agents. Cumulative incidences of HCC were analyzed after adjusting for competing mortality.

Because death usually results from underlying comorbidities that can also impact HCC risk, its occurrence leads to informative censoring in calculating HCC incidence. Therefore, death before HCC development was considered a competing risk event on survival analysis. Death-adjusted cumulative incidences in competing risk data ratios were analyzed using modified Kaplan-Meier method and Gray method.^{28,29} The R package "cmprsk" (<http://cran.r-project.org/web/packages/cmprsk/index.html>) was used in the competing risk analyses.

Multivariable Analyses

To determine whether nucleos(t)ide analogue use was independently associated with reduced risk of HCC, the Cox proportional hazards model was developed. Because the case numbers in the present study were large enough, all the parameters defined a priori were used in the model to exclude as many confounders as possible.

To examine potential heterogeneity of treatment effect in relation to confounders, we performed interaction analyses by adding interaction terms between treatment and potential

confounding factors, including age, sex, cirrhosis, liver decompensation, comorbidities, use of statins, use of nonsteroidal anti-inflammatory drugs, and use of metformin in the multivariable analyses. Because the sample size is quite large in the present study, we defined the α level at .05 as significantly different for the interaction instead of using an α level at .10. Once a factor is found significantly interacted with treatment, multivariable stratified analyses were conducted to examine the associations of nucleos(t)ide use and risk of HCC in CHB patients with these factors.

Sensitivity Analyses

Nucleos(t)ide analogues have been covered under the NHI program since October 1, 2003, midway through the follow-up period. Therefore, the follow-up durations in the 2 treatment groups differed significantly. The impact of reimbursement of antivirals in 2003 can induce several other unmeasured factors, such as different screening policy, different accessibility of health care, and different health consciousness of patients, etc. To address the differential follow-up in the 2 treatment groups, we conducted sensitivity analyses with fixed duration by limiting the index date of follow-up to between October 1, 2003 and September 30, 2005 and the follow-up duration to 5 years.

To examine the impact of other potential unmeasured confounders on the estimated treatment effect, we performed sensitivity analyses with add-on of an unmeasured confounder according to the method of Lin et al³⁰ using the R package "obsSens." In the sensitivity analyses, we added another hypothetical unmeasured confounder with a similar favorable protective effect as our antiviral agents. Then we examined how this added unmeasured factor confounded our observations with different prevalences in the treated and untreated groups.

Statistical Analysis

The demographic characteristics of the treated and untreated patients were compared using the χ^2 test and Student *t* test. The number of patients who needed to be treated for 1 less HCC occurrence was defined as NNT and calculated by the inverse of the absolute risk reduction. The confidence intervals of NNT were calculated according to the method of Altman et al in a survival setting.³¹ All data management was performed using SAS 9.1 software (SAS Institute, Cary, NC). Cumulative incidences were analyzed using the survival and EpiTools packages of R.

Results

Demographic Data

Between 1997 and 2010, a total of 199,451 patients were diagnosed with CHB. Among them, 81,823 patients had used nucleos(t)ide analogues or hepatoprotectants for at least 90 days. Patients with hepatitis C, other hepatitis, human immunodeficiency virus, and those with malignancies before the use of nucleos(t)ide analogues or hepatoprotectants were excluded. Among the remaining 72,458 patients, 47,611 patients were in the untreated group and 24,847 patients were in the treated group. These 2 groups had significant differences in their demographic data (Supplementary Table 1). To estimate the probability of

receiving nucleos(t)ide analogue therapy, age, sex, comorbidities (in separate terms for individual comorbidities), use of statins, use of nonsteroidal anti-inflammatory drugs, and use of metformin were used to calculate propensity scores. The original propensity scores of the untreated patients (mean, 0.29) were significantly lower when compared with the treated cohort (mean, 0.44) ($P < .001$). We used propensity score to match 1 patient in the treated cohort with 1 patient in the untreated cohort. The histograms of propensity score before and after matching are shown in Supplementary Figure 1. Before matching, the untreated patients had significantly lower propensity scores (*black bars* in the histogram) compared with the treated patients (*gray bars* in the histogram) (Supplementary Figure 1A). After matching, the 2 groups had comparable distributions of propensity scores (Supplementary Figure 1B). Finally, we recruited 21,595 patients into the treated cohort and 21,595 patients into the untreated cohort (Figure 1). In the treated cohort, a total of 19,063 patients received only 1 nucleos(t)ide analogue, including 12,938 patients who received lamivudine, 5748 patients who received entecavir, and 377 patients who received telbivudine. The remaining 2532 patients received more than 1 nucleos(t)ide analogue.

As shown in Table 1, the mean age of these 2 cohorts was 43.5 years and about three fourths of the patients were male. Mean follow-up durations for the treated

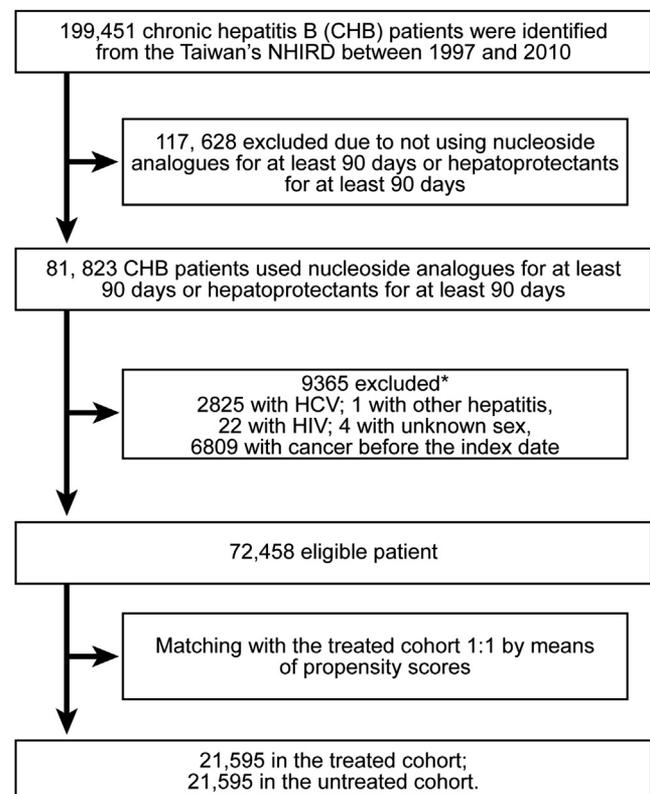


Figure 1. Study patient selection flow diagram. *A case can be excluded due to more than one criterion. Therefore, the total excluded cases in each step can outnumber the sum of case numbers excluded by individual criteria. HIV, human immunodeficiency virus.

and untreated cohorts were 3.46 years and 5.24 years, respectively. Mean sonography and α -fetoprotein screening frequencies (number per year) were 2.08 and 2.88 for the treated cohort and 1.90 and 1.81 for the untreated cohorts, respectively. Mean duration of nucleos(t)ide analogue use in the treated cohort was 1.44 years. Mean duration of hepatoprotectant use in the untreated cohort was 1.24 years. The treated cohort had a significantly lower incidence of

Table 1. Baseline Demographic Characteristics and Outcomes of Study Cohorts

Demographic characteristics and outcomes	Untreated ^a (n = 21,595)	Treated ^b (n = 21,595)	P value ^c
Age, y, mean \pm SD ^d	43.52 \pm 13.38	43.53 \pm 13.61	.935
Sex, n (%)			.001
Female	4982 (23.1)	5281 (24.5)	
Male	16,613 (76.9)	16,314 (75.5)	
Follow-up years ^e			
Mean \pm SD	5.24 \pm 2.17	3.46 \pm 2.2	<.001
Median (IQR)	6.51 (3.54–7.00)	3.34 (1.40–5.50)	<.001
Sonography screening (n/year)			
Mean \pm SD	1.90 \pm 1.58	2.08 \pm 1.61	<.001
Median (IQR)	1.54 (0.83–2.48)	1.85 (1.16–2.63)	<.001
AFP screening (n/year)			
Mean \pm SD	1.81 \pm 1.69	2.88 \pm 2.29	<.001
Median (IQR)	1.38 (0.56–2.55)	2.41 (1.38–3.84)	<.001
Nucleos(t)ide analogue therapy duration, y			
Mean \pm SD		1.44 \pm 0.74	
Median (IQR)		1.42 (1.02–1.68)	
Hepatoprotective agents, y			
Mean \pm SD	1.24 \pm 1.28	0.78 \pm 1.14	<.001
Median (IQR)	0.77 (0.43–1.55)	0.36 (0.08–0.98)	<.001
Concomitant drug users, ^f n (%)			
Statin	1413 (6.5)	1474 (6.8)	.248
NSAIDs or aspirin	11,996 (55.5)	11,903 (55.1)	.373
Metformin	1880 (8.7)	2000 (9.3)	.045
Major coexisting diseases, n (%)			
Cirrhosis	3016 (14.0)	2847 (13.2)	.018
Liver decompensation	1646 (7.6)	1695 (7.8)	.387
Hypertension	1827 (8.5)	1893 (8.8)	.265
Diabetes	1574 (7.3)	1590 (7.4)	.782
COPD	486 (2.3)	462 (2.1)	.450
Acute coronary syndrome	674 (3.1)	654 (3.0)	.596
Cerebral vascular disease	506 (2.3)	499 (2.3)	.848
Renal failure	389 (1.8)	414 (1.9)	.393
Hypercholesterolemia	206 (1.0)	201 (0.9)	.842
Charlson's score			
Mean \pm SD	0.80 \pm 1.53	0.79 \pm 1.52	.308
Median (IQR)	0 (0–1)	0 (0–1)	.165
Propensity score ^g			
Mean \pm SD	0.42 \pm 0.16	0.42 \pm 0.16	.476
Median (IQR)	0.42 (0.3–0.53)	0.42 (0.3–0.53)	.397
Events, n (%)			
HCC occurrence	4454 (20.6)	992 (4.6)	<.001
Death before HCC occurrence	2556 (11.8)	1036 (4.8)	<.001
Overall death	4778 (22.1)	1406 (6.5)	<.001

AFP, α -fetoprotein; COPD, chronic obstructive pulmonary disease; HCC, hepatocellular carcinoma; IQR, interquartile range; NSAIDs, nonsteroidal anti-inflammatory drugs.

^aUntreated, not receiving nucleos(t)ide analogues.

^bTreated, receiving nucleos(t)ide analogues.

^cP values were compared using χ^2 test and Student *t* test.

^dAge is treated as a continuous variable.

^eFollow-up is defined as the time of nucleos(t)ide analogue or hepatoprotective treatment.

^fDrug users indicate patients using drugs at least 1 day per month on average.

^gAge, sex, acute coronary syndrome, cerebral vascular diseases, COPD, diabetes, cirrhosis, HCC, liver decompensation, renal failure, hypertension, hypercholesterolemia, use of statins, use of NSAIDs or aspirin or coxibs, and use of metformin were included in the propensity score calculation.

HCC (n = 992 [4.6%]) when compared with the untreated cohort (n = 4454 [20.6%]). Competing mortality (death before development of HCC) was significantly lower in the treated cohort (n = 1036 [4.8%]) than in the untreated cohort (n = 2256 [11.8%]). Overall mortality in the treated cohort (n = 1406 [6.5%]) was also significantly lower than in the untreated cohort (n = 4778 [22.1%]).

Seven-Year Cumulative Incidences of Hepatocellular Carcinoma for Treated and Untreated Cohorts

Cumulative incidences of HCC after adjustment for competing mortality are shown in Figure 2. Patients in the treated cohort were associated with a significantly lower risk of HCC (7-year cumulative incidence: 7.32%; 95% confidence interval [CI]: 6.77%–7.87%) than those in the untreated cohort (22.70%; 95% CI: 22.11%–23.30%) ($P < .001$). On average, the annual incidences of HCC in treated and untreated cohorts were 1.05% and 3.24%, respectively. The unadjusted NNT associated with 1 less HCC development within 7 years was 7 (95% CI: 6.2–6.9). This suggests that use of nucleos(t)ide analogues in 7 CHB patients is associated with 1 less HCC development within 7 years.

Multivariable Analysis

Without controlling for other factors, nucleos(t)ide analogue treatment was associated with reduced risk of HCC development (hazard ratio [HR] = 0.34; 95% CI: 0.32–0.37; $P < .001$). After adjusting for competing mortality and other confounders, we found that nucleos(t)ide analogues treatment is associated with a significantly lower risk of HCC (HR = 0.37; 95% CI: 0.34–0.39; $P < .001$). Older age, male sex, and liver cirrhosis were found to be risk factors for

increased HCC risk. Patients with comorbidities including liver decompensation, hypertension, chronic obstructive pulmonary disease, acute coronary syndrome, and cerebral vascular diseases were found to be associated with reduced risk of HCC due to higher competing mortality. Use of statin and use of nonsteroidal anti-inflammatory drugs or aspirin were associated with significantly lower risk of HCC in CHB patients (Table 2).

To examine whether significant heterogeneity of treatment effect exists in relation to age, sex, cirrhosis, liver decompensation, diabetes, and other potential confounders, we added interaction terms to the multivariable analyses (Supplementary Table 2). On the interaction analysis, we found statistically significant interactions between nucleos(t)ide use and age, sex, liver cirrhosis, and diabetes. Because the interactions are statistically significant, we cannot interpret the main effect of treatment. Instead, we examined the effect of treatment within each level of the factors, including age, sex, liver cirrhosis, and diabetes. In Figure 3, we conducted multivariable subgroup analyses. We found that the treated cohort was associated with a reduced risk of HCC in all subgroups. The beneficial effect of nucleos(t)ide analogues was especially significant among younger patients (younger than 40 years old: HR = 0.13; 40–50 years old: HR = 0.30; 50 years or older: HR = 0.49), patients without cirrhosis (noncirrhosis vs cirrhosis: HR = 0.27; vs HR = 0.72), and patients without diabetes (non-diabetes vs diabetes: HR = 0.34 vs HR = 0.69).

CLINICAL LIVER

Table 2. Multivariable Cox Proportional Hazards Model Analysis for Risk of Hepatocellular Carcinoma Occurrence After Adjusting for Competing Mortality

	Hazard ratio ^a	95% CI	P value
Treated vs untreated	0.37	0.34–0.39	<.001
Age, per each incremental year	1.06	1.06–1.06	<.001
Male	1.52	1.42–1.63	<.001
Cirrhosis	1.92	1.77–2.08	<.001
Liver decompensation	0.87	0.78–0.97	.013
Hypertension	0.81	0.72–0.90	<.001
Diabetes	1.05	0.93–1.17	.450
COPD	0.71	0.59–0.87	.001
ACS	0.66	0.55–0.79	<.001
CVA	0.50	0.40–0.64	<.001
Renal failure	0.84	0.68–1.04	.110
Hypercholesterolemia	0.87	0.61–1.23	.430
Statin use	0.55	0.47–0.63	<.001
NSAIDs or aspirin use	0.62	0.58–0.65	<.001
Metformin use	0.97	0.879–1.073	.570

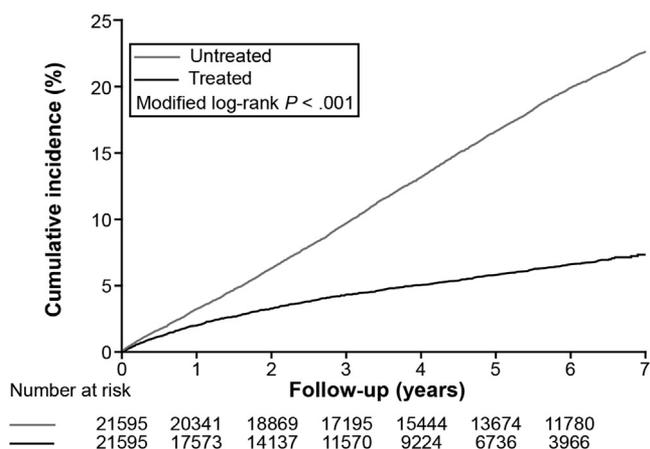


Figure 2. Cumulative incidences of HCC after adjustment for competing mortality. Calculation and comparison of cumulative incidences in competing risk data ratios were conducted using modified Kaplan–Meier method and Gray’s method. Patients who developed HCC during the first 3 months were excluded. Untreated, CHB patients not receiving nucleos(t)ide analogues; Treated, CHB patients receiving nucleos(t)ide analogues.

ACS, acute coronary syndrome; CI, confidence interval; COPD, chronic obstructive pulmonary disease; CVA, cerebral vascular diseases; NSAID, nonsteroidal anti-inflammatory drug.

^aAdjusted for covariate factors, including age, sex, comorbidities, use of statins, use of NSAIDs or aspirin, and use of metformin.

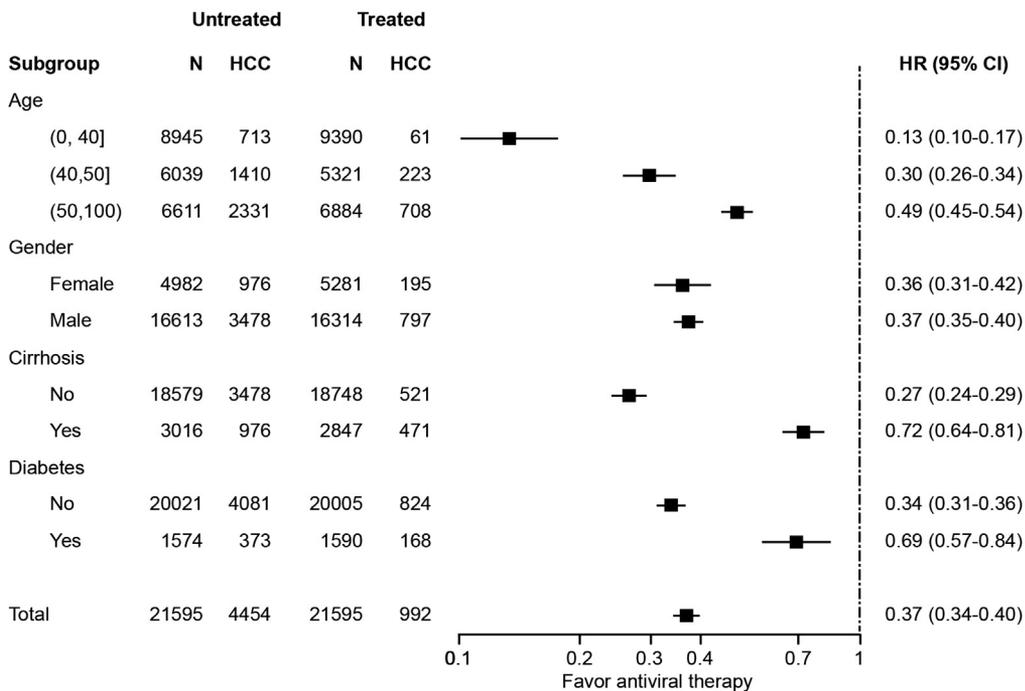


Figure 3. Multivariable stratified analyses. Among chronic CHB patients, nucleoside analogue use (treated cohort) is associated with reduced risk of HCC development in all subgroups. All *P* values were significant.

Sensitivity Analysis

In the sensitivity analysis with fixed duration, we only identified CHB patients between October 1, 2003 and September 30, 2005 because of 3 reasons. First, nucleos(t)ide analogues were covered under the NHI program since October 1, 2003. Second, we wanted to follow these patients up to 5 years, until the end of 2010. Third, these patients used nucleos(t)ide analogues or hepatoprotectants for at least 90 days. Propensity scores were used to match each treated patient with 1 untreated patient. Finally, we identified 4545 patients in the treated cohort and 4545 patients in the untreated cohort. The demographic characteristics and outcomes are shown in [Supplementary Table 3](#). Patients in the treated cohort were associated with significantly lower risk of HCC development (5-year cumulative incidence: 6.62%; 95% CI: 5.90%–7.35%) than those in the untreated cohort (19.08%; 95% CI: 17.93%–20.22%) ($P < .001$) ([Supplementary Figure 2](#)). On multivariable analysis, the treated cohort was associated with reduced risk of HCC development (adjusted HR = 0.31; 95% CI: 0.27–0.53; $P < .001$) ([Supplementary Table 4](#)).

In [Supplementary Figure 3](#), we used sensitivity analysis to examine the trend of estimates of the treated hazard on covariate-adjusted Cox model with add-on of an unmeasured confounder with relative hazard of 0.3. When all the subjects in untreated group have the add-on unmeasured confounder (prevalence of the confounder in the untreated group is 1.0) and none of subjects in treated group has this unmeasured confounder (prevalence of the confounder in the treated group is 0), then the impact of antiviral therapy would be beneficial (HR = 0.1, the *bottom line* in [Supplementary Figure 3](#)). On the contrary, when

none of subjects in the untreated group has the add-on unmeasured confounder ($P_0 = 0$) and all subjects in treated group have this confounder; then the impact of antiviral therapy would not be protective (HR = 1.2, the *top line* in [Figure 3](#)). In most situations, patients who received nucleos(t)ide analogues had lower risk of HCC occurrence relative to those who did not, even if a favorable unmeasured confounder exists.

Discussion

This population-based cohort study demonstrated that use of nucleos(t)ide analogues is associated with reduced long-term risk of HCC in CHB patients. After adjusting for death as the competing cause of risk and for multiple confounding factors, we found that use of nucleos(t)ide analogues is associated with an adjusted HR of 0.37 for HCC occurrence in CHB patients. The association between nucleos(t)ide analogues use and lower risk of HCC was found in all subgroups of CHB patients, especially in younger patients, patients without liver cirrhosis, patients without liver decompensation, and patients without diabetes. We also validated our observations by sensitivity analyses.

In the present study, the association between use of nucleos(t)ide analogues and risk of HCC in CHB patients diminished with age. The HRs associated with use of nucleos(t)ide analogues were 0.13, 0.30, and 0.49 for patients younger than 40 years old, aged between age 40 and 50 years, and older than 50 years, respectively. Several reasons might explain this observation. First, this interaction resulted from the rising probability of HCC in CHB patients with advanced age. In the REVEAL study (Risk

Evaluation of Viral Load Elevation and Associated Liver Disease/Cancer), which investigated the natural history of CHB patients in Taiwan, Chen et al found that the risk of HCC in CHB patients remains low before age 40 years, starts to rise in the 40s, and significantly increases after age 50 years.⁷ The rising probability of HCC with advanced age might be reflected in the absolute risk reduction by use of nucleos(t)ide analogues in different age groups. The 7-year HCC absolute risk reductions for patients younger than 40 years, aged 40–50 years, and older than 50 years were 7.92%, 18.23%, and 18.68%, respectively. Second, older patients might have more comorbidities, which leads to higher risk of competing mortality. In the present study, we found that patients with hypertension, chronic obstructive pulmonary disease, acute coronary syndrome, and cerebral vascular diseases have higher competing mortality and lower HCC risk after adjusting for competing mortality. Third, the starting time for nucleos(t)ide analogue use might be too late to rescue the carcinogenesis of HCC. However, we need more evidence to support this hypothesis.

For noncirrhotic CHB patients, mean annual incidences of HCC were 0.68% and 2.97% for the treated and the untreated cohorts, respectively. Adjusted HR was 0.25 for the use of nucleos(t)ide analogues. Our observations of the treated group were comparable with those of previous reports, but the annual HCC incidence in the untreated noncirrhotic patients in the present study was higher than in previous studies.^{13,20,32} In a meta-analysis based on 5 studies comparing patients treated with nucleos(t)ide analogues with controls, Sung et al reported that the risk of HCC after nucleos(t)ide analogue treatment is reduced by 78%.¹⁹ In another systematic review of 21 studies by Papatheodoridis et al, HCC developed in 2.8% and 6.4% of nucleos(t)ide analogue-treated and untreated CHB patients, respectively, during a 46-month period.²⁰ In a retrospective cohort study based on 377 CHB patients (17% with cirrhosis), annual HCC incidences were 0.4% for nucleos(t)ide analogue-treated group and 2.5% for control group, respectively.³³ The higher annual HCC incidence in our untreated noncirrhotic patients might be due to the requirement for higher baseline aminotransferase levels to obtain reimbursement for hepatoprotective agents.

For cirrhotic CHB patients, the mean annual incidences of HCC in the present study were 3.90% and 4.94% for treated and untreated cohorts, respectively. The adjusted HR was 0.72 for the use of nucleos(t)ide analogue. The annual incidence of our treated cohort was slightly higher than that of previous studies. In a retrospective cohort study based on CHB patients with cirrhosis, the annual incidences of HCC in nucleos(t)ide analogue treated and untreated groups were 1.02% and 6.0%, respectively.³⁴ In the CALM study, a randomized trial based on CHB patients with cirrhosis or advanced fibrosis, nucleos(t)ide analogue use was found to reduce risk of HCC development during a median duration of 32.4 months of therapy (3.9% vs 7.4%; $P = .047$).¹² A possible explanation for the higher annual incidences in the treated cohort in the present study is the strict NHI regulations regarding nucleos(t)ide analogue reimbursement.¹⁸ Only high-risk populations, including

patients with higher baseline HBV viral load, higher ALT level, or higher prevalence of liver decompensation are eligible for reimbursement. These higher-risk populations can contribute to the higher annual incidences.

The chemopreventive effect of nucleos(t)ide analogue therapy in the present study was significantly higher in nondiabetic patients when compared with diabetic patients (adjusted HR = 0.34 vs 0.69). In our recent nationwide case-control study, we found that diabetes is independently associated with increased risk of HCC development (odds ratio = 2.25).³⁵ In the United States, diabetes has also been found to be associated with 2- to 3-fold increase in the risk of HCC, regardless of other HCC risk factors, such as viral hepatitis.^{36,37} Several factors can explain the association between HCC and diabetes, such as increased risk for nonalcoholic fatty liver disease and nonalcoholic steatohepatitis in diabetic patients. Metformin has been found to be associated with a decreased risk of HCC in diabetes patients, via inhibition of hepatoma cell proliferation and induction of cell cycle arrest at G₀/G₁ phase.³⁵ In the era of nucleos(t)ide analogue therapy, more studies are needed to investigate the role of diabetes in the carcinogenesis of HCC to further decrease HCC incidence.

Recently, several HCC score calculators have been introduced.^{38–40} Unfortunately, we did not have information on baseline ALT and HBV DNA levels, which have been shown to be associated with HCC risk in these HCC score calculators. Given that reimbursement for nucleos(t)ide analogues requires twice-elevated aminotransferase and higher HBV DNA levels (>2000 IU/mL), and reimbursement for hepatoprotective agents (control group) requires only elevated aminotransferase level (ALT $\geq 1\times$), the higher baseline HBV DNA levels and ALT levels in the treated cohort might have led to a more conservative estimation of the protective effect of nucleos(t)ide analogues.

In the present study, we used many methods to prevent potential confounders. However, some unmeasurable bias can still exist. Propensity score matching was used to select comparable controls to imitate a randomized clinical trial. Although we used all potential confounders in the model to create propensity score, there were some significant differences in the distributions, such as for sex and liver cirrhosis. The large sample size in the present study might be the reason for the statistical significance. For examples, the differences between 24.5% female in the treated cohort vs 23.1% female in the untreated cohort, and the differences of 13.2% vs 14.0% cirrhosis in the treated and untreated cohorts are statistically significant, but might be not clinically significant. Reimbursement for nucleos(t)ide analogues began in 2003, midway through the follow-up period, which might also have confounded our observations. However, protective role of nucleos(t)ide analogue uses in HCC risk was found on sensitivity analyses with fixed duration.

In conclusion, nucleos(t)ide analogue use is associated with reduced risk of HCC in CHB patients. Age, sex, liver cirrhosis, and diabetes mellitus modify this association. More studies are needed to explore the wider use of nucleos(t)ide analogues for prolonged periods to decrease the incidences of HCC.

Supplementary Material

Note: To access the supplementary material accompanying this article, visit the online version of *Gastroenterology* at www.gastrojournal.org, and at <http://dx.doi.org/10.1053/j.gastro.2014.03.048>.

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Conflicts of interest

The authors disclose no conflicts.

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Supplementary Table 1. Baseline Demographic Characteristics and Outcomes of Study Cohorts Before Propensity Score Matching

	Untreated ^a (n = 47,611)	Treated ^b (n = 24,847)	P value ^c
Age, y, mean ± SD ^d	51.62 ± 14.34	41.95 ± 13.70	<.001
Sex, n (%)			<.001
Female	13,021 (27.3)	6087 (24.5)	
Male	34,590 (72.7)	18760 (75.5)	
Follow-up years, mean ± SD ^e			
Mean ± SD	5.18 ± 2.17	3.52 ± 2.2	<.001
Median (IQR)	6.32 (3.48–7.00)	3.44 (1.47–5.55)	<.001
Nucleos(t)ide analogue therapy duration, y			
Mean ± SD		1.42 ± 0.73	
Median (IQR)		1.41 (1.00–1.63)	
Hepatoprotective agents, y			
Mean ± SD	1.41 ± 1.44	0.74 ± 1.11	<.001
Median (IQR)	0.87 (0.46–1.80)	0.34 (0.07–0.93)	<.001
Concomitant drug users ^f			
Statin	6260 (13.1)	1529 (6.2)	<.001
NSAIDs or aspirin	34,615 (72.7)	12,395 (49.9)	<.001
Metformin	11,004 (23.1)	2017 (8.1)	<.001
Major coexisting diseases			
Cirrhosis	5179 (10.9)	3172 (12.8)	<.001
Liver decompensation	2087 (4.4)	2121 (8.5)	<.001
Hypertension	5133 (10.8)	2055 (8.3)	<.001
Diabetes	4415 (9.3)	1713 (6.9)	<.001
COPD	1829 (3.8)	491 (2.0)	<.001
Acute coronary syndrome	2214 (4.7)	676 (2.7)	<.001
Cerebral vascular disease	2535 (4.9)	506 (2.0)	<.001
Renal failure	676 (1.4)	496 (2.0)	<.001
Hypercholesterolemia	747 (1.6)	209 (0.8)	<.001
Charlson's score			
Mean ± SD	0.76 ± 1.43	0.77 ± 1.51	.294
Median (IQR)	0 (0–1)	0 (0–1)	<.001
Propensity score ^g			
Mean ± SD	0.29 ± 0.17	0.44 ± 0.17	<.001
Median (IQR)	0.25 (0.15–0.41)	0.45 (0.32–0.58)	<.001
Events			
HCC occurrence	11,574 (24.3)	1059 (4.3)	<.001
Death before HCC occurrence	5972 (12.5)	1185 (4.8)	
Overall death	11,865 (24.9)	1583 (6.4)	<.001

COPD, chronic obstructive pulmonary disease; IQR, interquartile range; NSAID, nonsteroidal anti-inflammatory drug.

^aUntreated, not receiving nucleos(t)ide analogues.

^bTreated, receiving nucleos(t)ide analogues.

^cP values were compared using the χ^2 test and Student *t* test.

^dAge is treated as a continuous variable.

^eFollow-up is defined as the time of nucleos(t)ide analogue or hepatoprotective treatment.

^fDrug users indicate patients using drugs at least 1 day per month on average.

^gAge, sex, acute coronary syndrome, cerebral vascular diseases, COPD, diabetes, cirrhosis, liver decompensation, renal failure, hypertension, hypercholesterolemia, use of statins, use of NSAIDs or aspirin or coxibs, and use of metformin were included in the propensity score calculation.

Supplementary Table 2. Interaction Analysis: Multivariable Cox Proportional Hazards Model Analysis for Risk of Hepatocellular Carcinoma Occurrence After Adding Interactions Between Therapy and Subgroup Factors

	Hazard ratio	95% CI	P value
Covariates			
Treated vs untreated	0.08	0.06–0.11	<.001
Age, per year	1.05	1.05–1.06	<.001
Male sex	1.48	1.34–1.55	<.001
Cirrhosis	1.60	1.46–1.76	<.001
Liver decompensation	0.84	0.74–0.95	.005
Hypertension	0.81	0.72–0.92	.001
Diabetes	0.99	0.87–1.13	.900
COPD	0.71	0.57–0.88	.002
Acute coronary syndrome	0.67	0.55–0.83	<.001
Cerebral vascular disease	0.41	0.30–0.55	<.001
Renal failure	0.89	0.70–1.13	.330
Hypercholesterolemia	0.83	0.55–1.24	.360
Statin use	0.54	0.46–0.64	<.001
NSAIDs or aspirin use	0.60	0.57–0.64	<.001
Metformin use	0.92	0.82–1.03	.150
Interactions with treatment			
Age (per year)*treatment	1.02	1.01–1.02	<.001
Male*treatment	1.31	1.09–1.58	.004
Cirrhosis*treatment	2.06	1.69–2.50	<.001
Liver Decompensation*treatment	1.21	0.96–1.52	.110
Hypertension*treatment	0.98	0.76–1.27	.880
Diabetes*treatment	1.33	1.02–1.73	.034
COPD*treatment	0.94	0.61–1.44	.760
Acute coronary syndrome*treatment	0.91	0.58–1.43	.690
Cerebral vascular disease*treatment	1.56	0.93–2.60	.090
Renal failure*treatment	0.85	0.50–1.43	.530
Hypercholesterolemia*treatment	1.19	0.52–2.70	.690
Statin use*treatment	0.92	0.65–1.31	.650
NSAIDs or aspirin use*treatment	1.09	0.94–1.26	.270
Metformin use*treatment	1.12	0.87–1.43	.380

COPD, chronic obstructive pulmonary disease; NSAIDs, nonsteroidal anti-inflammatory drugs.

Supplementary Table 3. Sensitivity Analysis With Fixed Duration: Baseline Demographic Characteristics and Outcomes

	Untreated ^a (n = 4545)	Treated ^b (n = 4545)	P value ^c
Age, y, mean ± SD ^d	44.81 ± 12.82	44.67 ± 12.99	.609
Sex			.660
Female	1117 (24.6)	1098 (24.2)	
Male	3428 (75.4)	3447 (75.5)	
Follow-up years ^e			
Mean ± SD	4.26 ± 1.37	4.63 ± 1.10	<.001
Median (IQR)	5.00 (4.1–5.00)	5.00 (5.00–5.00)	<.001
Nucleoside analogue therapy duration, y			
Mean ± SD		1.46 ± 0.84	
Median (IQR)		1.44 (1.02–1.53)	
Hepatoprotective agents, y			
Mean ± SD	1.09 ± 1.02	0.92 ± 1.25	<.001
Median (IQR)	0.71 (0.42–1.37)	0.47 (0.13–1.19)	<.001
Concomitant drug users ^f			
Statin	409 (9.0)	420 (9.2)	.716
NSAIDs or aspirin	2869 (63.1)	2851 (62.7)	.712
Metformin	529 (11.6)	548 (12.1)	.559
Major coexisting diseases			
Cirrhosis	517 (11.4)	517 (11.4)	1.000
Liver decompensation	269 (5.9)	284 (6.2)	.539
Hypertension	381 (8.4)	352 (7.7)	.281
Diabetes	323 (7.1)	305 (6.7)	.482
COPD	107 (2.4)	94 (2.1)	.392
Acute coronary syndrome	125 (2.8)	119 (2.6)	.746
Cerebral vascular disease	105 (2.3)	89 (2.0)	.276
Renal failure	82 (1.8)	79 (1.7)	.874
Hypercholesterolemia	53 (1.2)	45 (1.0)	.477
Charlson's score			
Mean ± SD	0.76 ± 1.45	0.70 ± 1.41	.031
Median (IQR)	0 (0–1)	0 (0–1)	0
Propensity score ^g			
Mean ± SD	0.54 ± 0.20	0.54 ± 0.20	.451
Median (IQR)	0.55 (0.41–0.70)	0.55 (0.41–0.70)	.368
Events			
HCC occurrence	867 (19.1)	301 (6.6)	<.001
Death before HCC occurrence	489 (10.8)	282 (6.2)	<.001
Overall death	884 (19.4)	405 (8.9)	<.001

NOTE. Because nucleoside analogues have been reimbursed under the NHI program since October 1, 2003, we conducted sensitivity analysis by limiting index date of follow-up to between October 1, 2003, and September 30, 2005 and limiting the follow-up duration to 5 years.

COPD, chronic obstructive pulmonary disease.

^aUntreated, not receiving nucleoside analogues.

^bTreated, receiving nucleoside analogues.

^cP values were compared using the χ^2 test and Student *t* test.

^dTreating age as a continuous variable.

^eFollow-up is defined as the period of nucleoside analogue or hepatoprotective treatment.

^fDrug users indicate patients using drugs for at least 1 day per month on average.

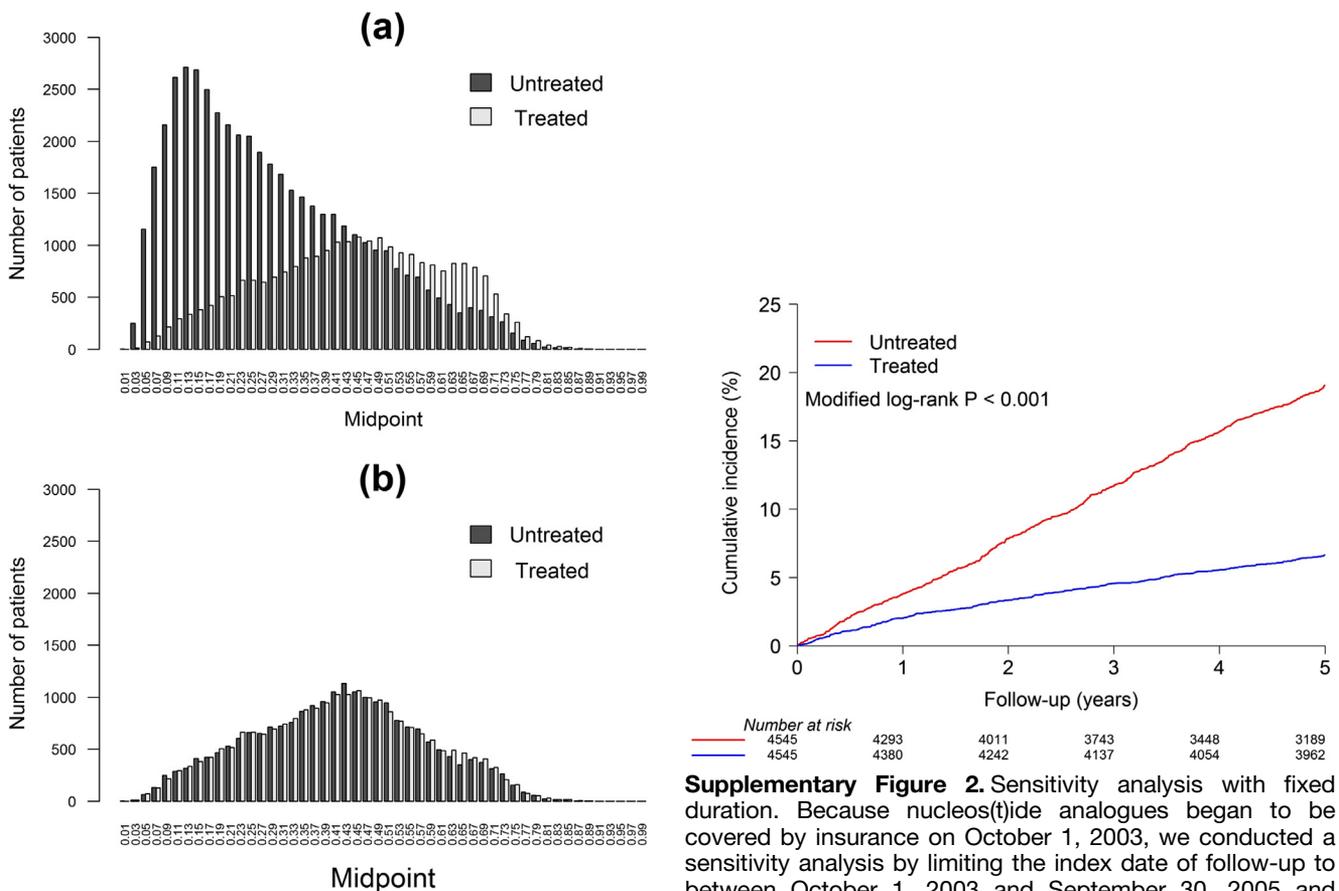
^gAge, sex, acute coronary syndrome, cerebral vascular diseases, COPD, diabetes, cirrhosis, liver decompensation, renal failure, hypertension, hypercholesterolemia, use of statins, use of NSAIDs, aspirin or coxibs, and use of metformin were included in the propensity score calculation.

Supplementary Table 4. Sensitivity Analysis With Fixed Duration: Multivariable Cox Proportional Hazards Model Analysis for Risk of Hepatocellular Carcinoma Occurrence

	Hazards ratio ^a	95% CI	P value
Treated vs untreated	0.31	0.27–0.35	<.001
Age, per each incremental year	1.06	1.06–1.07	<.001
Male	1.80	1.56–2.09	<.001
Cirrhosis	1.81	1.52–2.15	<.001
Liver decompensation	0.89	0.70–1.13	.340
Hypertension	0.74	0.58–0.94	.014
Diabetes	1.04	0.83–1.31	.740
COPD	0.65	0.42–1.00	.047
Acute coronary syndrome	0.76	0.51–1.15	.190
Cerebral vascular disease	0.54	0.34–0.85	.008
Renal failure	1.12	0.72–1.74	.620
Hypercholesterolemia	0.52	0.24–1.12	.094
Statin use	0.64	0.50–0.81	<.001
NSAIDs or aspirin use	0.61	0.54–0.69	<.001
Metformin use	1.19	0.99–1.43	.060

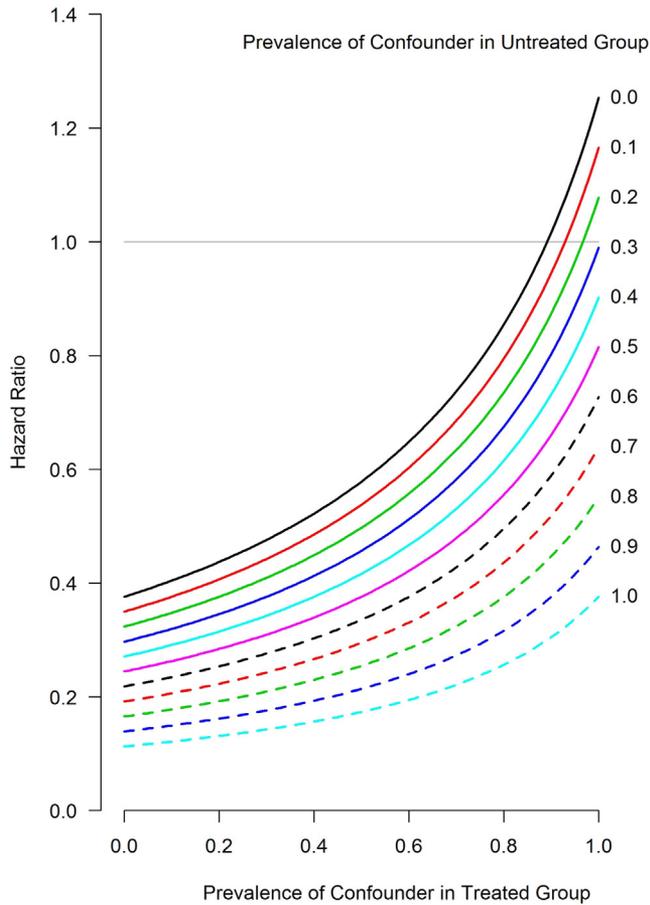
COPD, chronic obstructive pulmonary disease; NSAID, nonsteroidal anti-inflammatory drug.

^aAdjusted for covariate factors, including age, sex, comorbidities, use of statins, use of NSAIDs or aspirin and use of metformin.



Supplementary Figure 1. Histograms of propensity score before and after matching. (A) Histograms of propensity scores before matching. (B) Histogram of propensity scores after matching. Untreated, CHB patients not receiving nucleos(t)ide analogues; Treated, CHB patients receiving nucleos(t)ide analogues.

Supplementary Figure 2. Sensitivity analysis with fixed duration. Because nucleos(t)ide analogues began to be covered by insurance on October 1, 2003, we conducted a sensitivity analysis by limiting the index date of follow-up to between October 1, 2003 and September 30, 2005 and limiting the follow-up duration to 5 years. Calculation and comparison of cumulative incidences in competing risk data ratios were conducted using modified Kaplan–Meier method and Gray’s method. Untreated, CHB patients not receiving nucleos(t)ide analogues; Treated, CHB patients receiving nucleos(t)ide analogues.



Supplementary Figure 3. Sensitivity analysis with add-on of an unmeasured confounder. This figure displays the trend of estimates for the treated hazard on covariate-adjusted Cox proportional hazards model. For example, when all the subjects in the untreated group have the add-on unmeasured confounder (the prevalence of the confounder in the untreated group is 1.0) and none of subjects in treated group has this unmeasured confounder (the prevalence of the confounder in the treated group is 0.0), the impact of antiviral therapy would be beneficial (HR = 0.1, the *bottom line* in figure). On the contrary, when none of subjects in the untreated group has the add-on unmeasured confounder and all subjects in treated group have this confounder; then the impact of antiviral therapy would not be protective (HR = 1.2, the *top line* in the figure).